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2
3 **Variation in Rates of ADHD Diagnoses Across Public Schools in Arkansas**
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22 **Abstract**

23 The school context is important to diagnosing ADHD, but little is known about variation in rates
24 of ADHD diagnosis across schools. This study analyzed 87,383 kindergartners who were not
25 diagnosed with ADHD prior to kindergarten who attended at least one of 490 elementary schools
26 in Arkansas. The study examined five kindergarten cohorts from the 2015/16 to 2019/20
27 academic years, assessing school-level variation in ADHD diagnosis rates. Schools were
28 categorized by quintiles of their ADHD diagnosis rates. A mixed-effects logistic regression
29 model was used to examine individual-level factors associated with new ADHD
30 diagnosis. Diagnosis rates varied 570% across school quintiles. Schools with higher rates of
31 ADHD diagnoses served a higher percentage of African American children, children from lower-
32 income families, and children living in neighborhoods with lower Child Opportunity Index
33 (COI) scores. Results of an individual-level logistic regression analysis indicated children in
34 census tracts with very high or high social and economic COI scores had lower odds of having
35 new ADHD diagnoses compared to those from very low COI neighborhoods. Children who were
36 young-for-grade had higher odds of an ADHD diagnosis. Children who were in schools
37 providing universal free school meals had lower odds of an ADHD diagnosis, despite these
38 schools serving more children from families with lower incomes, a finding that warrants further
39 exploration. Because the rate of new ADHD diagnoses in kindergarten varies so substantially
40 across schools, future research is needed to better understand causes of school-level variability in
41 ADHD diagnosis.
42

1 **Introduction**

2 Attention-deficit/hyperactivity disorder (ADHD) is the most common childhood
3 neurodevelopmental disorder in the United States affecting ten percent of children of elementary
4 school age (1). School systems play an important role in the identification and diagnosis of
5 childhood ADHD. DSM-5 (2) criteria for diagnosis of ADHD require several symptoms to be
6 present in two or more settings (e.g., home and school). School staff, such as teachers, can be
7 influential in the referral of children for ADHD evaluation (5). Once referred, it is recommended
8 that ADHD evaluations for children include perspectives from teachers or other school staff
9 (3,4). Moreover, the classroom and school environment (e.g., level of structure, use of positive
10 behavioral interventions) are important for outcomes of children with ADHD (3,5).

11 Existing literature documents geographic variation in the prevalence of ADHD in
12 children in the United States (6), Norway (7), and Denmark (8). For example, Zgodic et al. found
13 that county-level prevalence of parent-reported childhood ADHD diagnoses in the United States
14 ranged from 11% or lower to 20% or higher (6). Other studies have examined variability in
15 ADHD diagnosis at the primary care practice level; these have observed substantial variability in
16 the ADHD diagnosis across primary care practices (9,10). However, little is known about the
17 extent to which incident ADHD diagnoses varies across schools. This is an important gap in the
18 literature because schools are an important context for the diagnosis and management of ADHD.
19 If rates of ADHD diagnosis are distributed unevenly across schools, this may signal potential
20 inequities in identification processes, school environments, and/or community-level factors that
21 contribute to ADHD prevalence rates. Given educational resources required to manage ADHD
22 are substantial (11), this also may have implications for the equitable distribution of resources to
23 support students with ADHD.

1 Geographic variation in the prevalence of ADHD has been attributed to a wide range of
2 factors, including those at the macro level such as education and health care system policy
3 (12,13); those related to clinicians such as availability of trained providers, clinicians' attitudes
4 and practice styles (14,15); and those related to population demographics and social determinants
5 of health such as poverty rates and access to community resources (16,17). Multifactorial genetic
6 and environmental risk factors are thought to interact to shape ADHD risk (18). Indeed, social
7 determinants of health within residential neighborhoods have been shown to be associated with
8 ADHD (35,36). These associations may reflect disparities in underlying risk as well as
9 disparities in identification, diagnosis and treatment access processes, and have been attributed to
10 a wide array of potential factors, including imbalanced community resources, exposure to stress,
11 environmental exposures, healthcare access, and sociocultural expectations (19–22). Because
12 schools tend to serve student populations grouped by geography, neighborhood-level factors may
13 lead to school-level variation in ADHD diagnosis. At the school level, there are several factors
14 that may be important for diagnosis and outcomes of children with ADHD and could lead to
15 school-level variation in diagnosis rates (23,24). First, some children access preventive and
16 diagnostic care directly through school-based health clinics (25). Second, In the United States,
17 children obtain up to 58% of their daily caloric intake from school meals (26), which are
18 required to conform to federal nutrition standards (27) and compare favorably to meals brought
19 from home in terms of overall nutritional quality (28–30), but vary across schools (31–33).
20 Based on recent systematic reviews, the evidence supporting a link between healthy diets
21 (34,35), sugar consumption (36) and ADHD diagnosis is not conclusive. However, this literature
22 primarily addresses individual dietary factors rather than school-level processes that facilitate
23 participation in child nutrition programs. Using difference-in-differences designs, recent studies

1 have shown that disciplinary infractions decline when schools adopt policies that facilitate
2 participation in child nutrition programs (37,38), which could impact diagnosis rates through
3 lower referrals for screening. Third, there is strong evidence for relative age effects in ADHD
4 diagnoses with younger children being more likely to receive an ADHD diagnosis than older
5 children in a grade cohort (39–42). Teacher ratings for ADHD-related symptoms often diverge
6 from parent ratings for children who are young-for-grade (40,42), highlighting the importance of
7 developmental maturity within the school context in establishing ADHD diagnoses. School
8 policies around discipline may also be relevant to ADHD diagnosis and management. Students
9 with ADHD are more likely to receive referrals and sanctions that include exclusionary
10 discipline (43), which is linked to a host of negative outcomes, such as involvement with the
11 justice system and risky behaviors (44). Recent evidence also suggests that attending a school
12 using a specific evidence-based positive behavior support framework (i.e., School-Wide Positive
13 Behavioral Interventions and Supports) reduces the likelihood of a student receiving
14 pharmacotherapeutic treatment for ADHD, especially for at-risk students in schools with high
15 implementation fidelity(45). Although not definitive, these results raise the possibility that
16 aspects of the school environment (e.g., level of proactive and positive behavioral support) may
17 influence symptom presentation for children with a genetic predisposition for ADHD (46–48).

18 Despite the myriad of factors that could lead to school-level variability in rates of ADHD
19 diagnosis (48), no existing studies have investigated school-level variation in children receiving
20 new ADHD diagnoses in their medical claims in the United States. Such studies are inherently
21 difficult due to federal and state privacy laws and regulations for both educational and medical
22 records, and the challenges in data access and linkage for epidemiologic studies. This study is the
23 first to report findings from such data, which allow us to simultaneously adjust for both

1 individual- and school-level factors. Participation in kindergarten is the norm across the United
2 States (49) and represents the first encounter with public-school systems for many children,
3 which makes it the ideal setting to examine school-level variation in ADHD diagnoses. The aims
4 of this cross-sectional study are to 1) describe school-level variation in rates of new ADHD
5 diagnoses among kindergarteners across public elementary schools; 2) examine child-level,
6 school-level, and neighborhood-level factors associated with kindergarten diagnosis rates.

7 **Methods**

8 *Design, Setting & Participants*

9 The study covered five first-time public-school kindergarten cohorts enrolled in Arkansas public
10 schools during the 2015/16 through 2019/20 academic years. This time frame was selected to
11 mitigate the influence of the COVID-19 pandemic, which caused disruptions to the delivery of
12 academic programs, potentially introducing confounding factors that could bias estimation. The
13 study involved the retrospective analysis of administrative datasets from the Arkansas
14 Department of Education (ADE) that were linked by subject identifiers to the Arkansas All-
15 Payer Claims Database (AR-APCD) through an honest data broker arrangement and were then
16 de-identified. Claims data, such as those in the AR-APCD, reflect the billing system used in US
17 healthcare, which is comprised of a fragmented, multi-payer system that encompasses private
18 insurers, employer-sponsored insurance plans, and public insurance (e.g. Medicaid). Claims
19 reflect billing information such as diagnostic codes (i.e., International Classification of Disease
20 (ICD) codes and procedure codes recorded and billed by clinicians and prescriptions billed to
21 insurance payers. Since the data available to the research team was de-identified, the study
22 received a not human subject research determination by the Institutional Review Board
23 (<<institution>> IRB# 275801).

1 Medical claims were evaluated over a 22-month period from September 1st of the year
2 prior to the year of kindergarten enrollment through June 30th of the year after enrollment. This
3 comprised a 12-month lookback period prior to entering kindergarten ending August 31st of the
4 year of kindergarten enrollment and a 10-month study period corresponding to the kindergarten
5 academic year. The validity of this approach is supported by earlier work showing positive
6 predictive values for ADHD cases identified in claims data are high when cross checked against
7 medical records (50).

8 We identified 116,350 first-time kindergartners during the study period (Figure 1).
9 Children were excluded for the following reasons: 1) The child did not have continuous
10 representation in the AR-APCD as defined by a coverage gap of no more than two consecutive
11 months during the 22-month lookback and study period (28,569 excluded). 2) To identify
12 incident ADHD, children with an ADHD diagnosis in the 12 months prior to September 1 of the
13 kindergarten year (4,939 excluded). 3) The child attended very small schools that had fewer than
14 25 kindergartners enrolled during the five-year period (398 excluded). These small schools were
15 excluded to avoid small-numbers bias in the school-level variations analysis. 4) The child was
16 missing study measures (37 excluded).

17 *Diagnoses*

18 We identified an ADHD diagnosis as either 1) a diagnostic code on two or more distinct dates or
19 2) a diagnostic code plus a prescription for a psychotropic stimulant (51). This approach will
20 identify cases where stimulants are not used to manage ADHD while reducing false positives
21 that could have occurred in inpatient settings, probable but not confirmed cases, or because of
22 “rule out” codes that may have been used prior to transitioning to the ICD-10-CM guidelines.
23 Diagnostic codes that were used to define ADHD were ICD-9-CM codes of 314.00, 314.01,

1 314.0 and ICD-10-CM codes of F90.0, F90.1, F90.2, F90.8, F90.9. Psychotropic stimulants were
2 defined as one or more prescriptions filled for amphetamine and mixed amphetamine salts,
3 dexmethylphenidate, dextroamphetamine, lisdexamfetamine, methamphetamine, and
4 methylphenidate (50). New cases diagnosed during kindergarten were defined as those where the
5 first diagnostic code or prescription occurred on or after September 1st of the year the child
6 entered kindergarten.

7 Co-occurring conditions (oppositional defiant disorder, other disruptive behavioral
8 disorders, autism spectrum disorder, major depressive disorder, anxiety disorders, stress
9 disorders, specific learning disorders, and tic disorders) were identified as one or more diagnostic
10 codes within the 22-month combined lookback and study periods. Other disruptive behavioral
11 disorders were comprised primarily by diagnostic code F91.9 (“conduct disorder unspecified”).
12 A complete list of diagnostic codes for these conditions is provided in eTable 1 in supplement 1.
13 To account for potential variation in diagnoses across medical providers, each child was assigned
14 to a provider who appeared most frequently in their history of medical claims for ambulatory
15 care during the 12-month lookback period. There were 3,579 distinct providers.

16 *Child-level factors*

17 Child-level measures were assessed using the ADE’s Statewide Information System (52). Based
18 on enrollment records, gender was coded as male or female and four classifications were used for
19 race/ethnicity: non-Hispanic White, non-Hispanic African American, non-Hispanic of another
20 race (including more than one race), or Hispanic. Indicators were assigned to children using a
21 school-based health center or receiving special education services any time during an academic
22 year. An indicator for school mobility was used to flag children who attended more than one
23 school during the year. For purposes of the analysis, the child flagged as mobile were assigned to

1 the school where they spent the most time during the academic year. Children were classified as
2 young-for-grade if they had birthdays in May, June, July, or August and were born exactly five
3 years before the year they entered kindergarten. An indicator was used to measure whether the
4 child received one or more disciplinary referrals during the school year. Lastly, to measure
5 socioeconomic status, indicators of eligibility for free or reduced-price meals, were used to
6 identify students from families below 130% and 185% of the federal poverty threshold,
7 respectively. For context, in 2018, the mid-year of our study, the federal poverty threshold for a
8 family of four was \$25,100 per annum and the median household income was \$63,179 per
9 annum (53).

10 *Neighborhood-level factors*

11 Childhood Opportunity Index (COI) scores were taken from the Child Opportunity Index 2.0
12 Database. Child opportunity was measured for the child's census tract using COI component
13 scores across three subdomains: (1) education (2) health and environment, and (3) economic and
14 social opportunities/resources, each ranging from 1 to 100, with higher numbers indicating
15 higher opportunity (54). Region indicators used by the Arkansas Center for Health Improvement
16 were used to capture broad differences in rurality and economic opportunity across the state
17 (eFigure 1 and eTable 2 in supplement 1).

18 *School-level factors*

19 Data on school meal policies were obtained from ADE's Child Nutrition Unit and the Arkansas
20 Hunger Relief Alliance. The study included two measures related to use of school meal-delivery
21 options that have been shown to increase access and participation to school nutrition. Indicators
22 were used to identify schools that provided universal free meals (UFM) to all enrolled children

1 and schools that provided “breakfast after the bell” (BAB), a meal delivery option that makes
2 breakfast available after the instructional component of the school day begins.

3 *Statistical methods*

4 Descriptive statistics were computed across all children. Variation across schools was assessed
5 by averaging study measures by school over the five academic years, which provided a rate of
6 new ADHD diagnoses by school. In the school-level analysis five-year averages were used to
7 ensure that a sufficiently large numbers of individuals were included to compute stable diagnosis
8 rates by school. Schools were then assigned to quintiles based on this rate. Trend tests were used
9 to assess significance in rates of ADHD diagnosis and the various child-level, neighborhood
10 level, and school-level measures described above across lowest and highest incidence schools.
11 The Cochran-Armitage test for trend was used for categorical variables and the Jonckheere–
12 Terpstra test for trend was used for continuous variables.

13 A mixed-effects logistic regression model with random effects for school, region, and
14 provider was used to examine the association between child-level ADHD diagnosis and the
15 child-level measures described above. Co-occurring conditions, the use of prescription
16 stimulants, and indicator for disciplinary infractions were excluded from this model because they
17 were likely to be codetermined with an ADHD diagnosis. The estimated model was used to
18 adjust school-level diagnoses rates. To align with the school-level analysis, conditional, model-
19 adjusted, individual-level predictions were averaged by school and then compared to the
20 unadjusted diagnosis rates.

21 All analyses were conducted using R version 4.4.0.

22 *Sensitivity Analysis*

1 Two sensitivity assessments were conducted. The first relaxed the 22-month continuous
2 enrollment exclusion requirement. This requirement was to reduce false-negatives, but could
3 impact sample composition. The second assessment involved estimation of two additional
4 mixed-effects logistic regression models, one excluding provider random effects and the other
5 excluding school random effects. These were to assess sensitivity to inclusion random effects
6 for provider and school.

7 **Results**

8 *Study Population*

9 After applying the exclusion criteria (Figure 1), the analytic sample contained 87,383
10 kindergartners. Supplemental material shows that those excluded were less likely to qualify for
11 free meals and were less likely to be racial and ethnic minorities (see eTable 3 in supplement 1).
12 Three out of four Arkansas kindergartners were from families with incomes below 185% of
13 poverty as indicated by eligibility for free or reduced-price school meals, with most
14 kindergartners (68%) falling below the free-meal threshold (130% poverty; Table 1).
15 Approximately 22% of kindergartners attended a school that offered UFM, 34% attended schools
16 that offered BAB and 8% of kindergartners received one or more disciplinary referrals.

17 *Variability in ADHD Diagnoses by School*

18 Across all schools, 5.2% of kindergartners were newly diagnosed with ADHD (Table 1). This
19 varied from 1.7% in the lowest quintile of schools to 10.0% in the highest quintile of schools, a
20 nearly six-fold difference (Table 2). Schools in the highest quintile served a higher percentage of
21 students who qualified for free school meals and who lived in neighborhoods with lower child
22 opportunity scores across all domains. African American children and children from the
23 Mississippi Delta region were more prevalent in the highest quintile of schools. A smaller

1 percentage of students in the highest quintile received care through a school-based health clinic
2 (1.6% versus 2.3%) while a higher percentage of students in the highest quintile were in schools
3 that provided UFM (31.5% versus 16.9%). The prevalence of co-occurring conditions, especially
4 ODD and CD, increased from the lowest to highest quintiles as did the use of stimulant
5 medications. A similar trend across quintiles was observed with respect to disciplinary referrals
6 and use of special education services.

7 *Associations with ADHD Diagnoses*

8 Table 3 presents the odds ratios (OR) and 95% confidence intervals (CIs) estimated from the
9 mixed-effect logistic regression model to predict individual ADHD diagnoses. Female children,
10 Hispanic children, non-Hispanic children of other races, children living in neighborhoods with
11 very high COI levels in the social and economic domain, and children attending schools offering
12 UFM had lower odds of being newly diagnosed with ADHD during kindergarten compared to
13 their respective reference groups. Conversely, students eligible for free or reduced-price meals,
14 those who transferred between schools, and those who were young-for-grade had significantly
15 higher odds of being diagnosed with ADHD compared to their reference groups.

16 Figure 2 shows unadjusted (panel A) and model-adjusted (panel B) rates of ADHD
17 diagnosis across the 490 schools. Figure 2 shows that the median observed ADHD diagnosis rate
18 shows a 4.9-times difference between the 10th and 90th percentiles whereas the median model-
19 adjusted ADHD diagnosis rate shows a 2.2-times difference between the 10th and 90th
20 percentiles.

21 *Sensitivity Analysis*

22 A comparison of findings including children without continuous coverage in the 12-month
23 lookback and 10-month study periods are reported in eFigure 2 and eTable 3-5 in Supplement 1.

1 As expected, diagnosis rates were lower without the continuous coverage requirement, but
2 primary findings align with those reported above. Associations from mixed-effects logistic
3 regression models were robust to the exclusion of provider effects or school effects. However,
4 the model included both school and provider effects had a lower Akaike Information Criterion,
5 indicating better model fit (eTable 6 in supplement 1).

6 **Discussion**

7 This study extends the literature on variability in ADHD diagnoses by examining school-level
8 variability in new medical claims-based ADHD diagnoses among kindergarteners using a novel
9 dataset with linked administrative educational and medical claims data for children in Arkansas.
10 We found substantial school-level variability: Schools in the highest quintile of new ADHD
11 diagnoses had a diagnosis rate nearly six times higher than those in the lowest quintile.
12 Compared to schools in the lowest quintile, schools in the highest quintile included a higher
13 proportion of African American children and children families eligible for free or reduced-price
14 school meals. We also observed several factors to be significantly associated with child-level
15 odds of receiving an ADHD diagnosis during kindergarten. After adjusting for these factors, the
16 difference in diagnostic rates across the 10th to 90th percentile of schools was reduced by half.
17 These results suggest potential factors to examine in future research designed to examine causal
18 relationships between school-level factors and rates of ADHD diagnoses.

19 The current findings extend prior work demonstrating variability in rates of childhood
20 ADHD diagnoses based on geography (6) and primary care practice (9) both within and outside
21 of the United States. Previous research in Norway has demonstrated geographical variability in
22 ADHD diagnoses with evidence that variability in this context may be driven by clinician
23 attitudes and local practice styles (10,55–57). However, our findings that mixed-effect models

1 were robust to inclusion of provider effects suggest that the school-level variation observed in
2 Arkansas is unlikely be explained solely by providers.

3 Although the associations observed here do not establish causal relationships, they do
4 have potential implications for policies that may promote equity in diagnoses. The school-level
5 variability in ADHD diagnoses observed here occurred in the context of school-level disparities
6 in resources and child academic and disciplinary outcomes observed in the literature (58–61).
7 Rates of ADHD diagnoses may interact with these school-level disparities in complex ways.
8 Indeed, schools with the highest ADHD diagnosis rates also had higher rates of disciplinary
9 referrals. Exclusionary discipline in elementary school contributes to an escalation of behavioral
10 issues, which may predispose the child to an ODD or CD diagnoses and/or interactions with the
11 juvenile justice system (62). Additionally, although UFM schools generally serve children from
12 neighborhoods facing economic disadvantage, we also found that attending a school with UFM
13 was associated with 21% lower adjusted odds of receiving an ADHD diagnosis, which is an
14 intriguing finding. There is a need for further probe and test mechanisms through which school
15 meal policy could impact diagnosis or outcomes for children with ADHD. Finally, our results
16 also replicate well-established findings about relative age effects (39–42,63) on ADHD
17 diagnosis, where children who are young for their grade level are more likely to be diagnosed,
18 which highlights the potential importance of ensuring that ADHD screening protocols are
19 cognizant of the child’s developmental stage relative to their classroom peers.

20 This study has several limitations. Although it is the first to link claims data to
21 educational records on school enrollment, it does not contain information on children who
22 attended kindergarten in private schools, who did not have insurance, or who were covered under
23 a self-insured employer plan that was not required to report to the AR-APCD. The use of claims

1 data poses limitations related to gaps in patient history, exclusion of cash payments, and inability
2 to assess adherence to recommendations and medications. In particular socioeconomic status can
3 affect medical coverage gaps and, in the data reported here, socioeconomic indicators were
4 significant predictors of an ADHD diagnosis. Because we required continuous coverage for
5 inclusion in the study, our data are likely over-represents children on public insurance and
6 children who were consistently on private insurance at the lower and upper ends of the
7 socioeconomic gradient, respectively. This study also did not capture related educational
8 classifications (e.g., Other Health Impairment) or ADHD-related services that were not reflected
9 in health insurance claims, which means that the overall rates of ADHD observed here may be an
10 underestimate. While the study draws data from a diverse statewide school system including both
11 urban and rural areas, findings from one state may not apply generally to other areas of the US or
12 elsewhere in the world, especially as healthcare access and educational policy vary across states
13 and countries. Finally, it is important to emphasize that the associations observed in this study
14 are observational and may be influenced by unmeasured confounding, so these associations
15 cannot be interpreted as causal.

16 **Conclusion**

17 Although many studies have observed variation in ADHD diagnoses, our contribution lies in
18 highlighting how this variation manifests at the school level in the United States and generates
19 new approaches for equitable diagnoses. Schools with the highest incidence of ADHD diagnoses
20 served children from less affluent families living in neighborhoods with lower economic
21 opportunity, which is broadly consistent with the idea of social determinants of health risk
22 factors that may interact with biological risk for ADHD (35). Across schools, higher incidences

- 1 aligned with a larger percentage of children being cited for one or more disciplinary infractions,
- 2 highlighting the importance of initiatives that prevent escalation of behavioral problems.

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Figure 1. Subject Attrition diagram

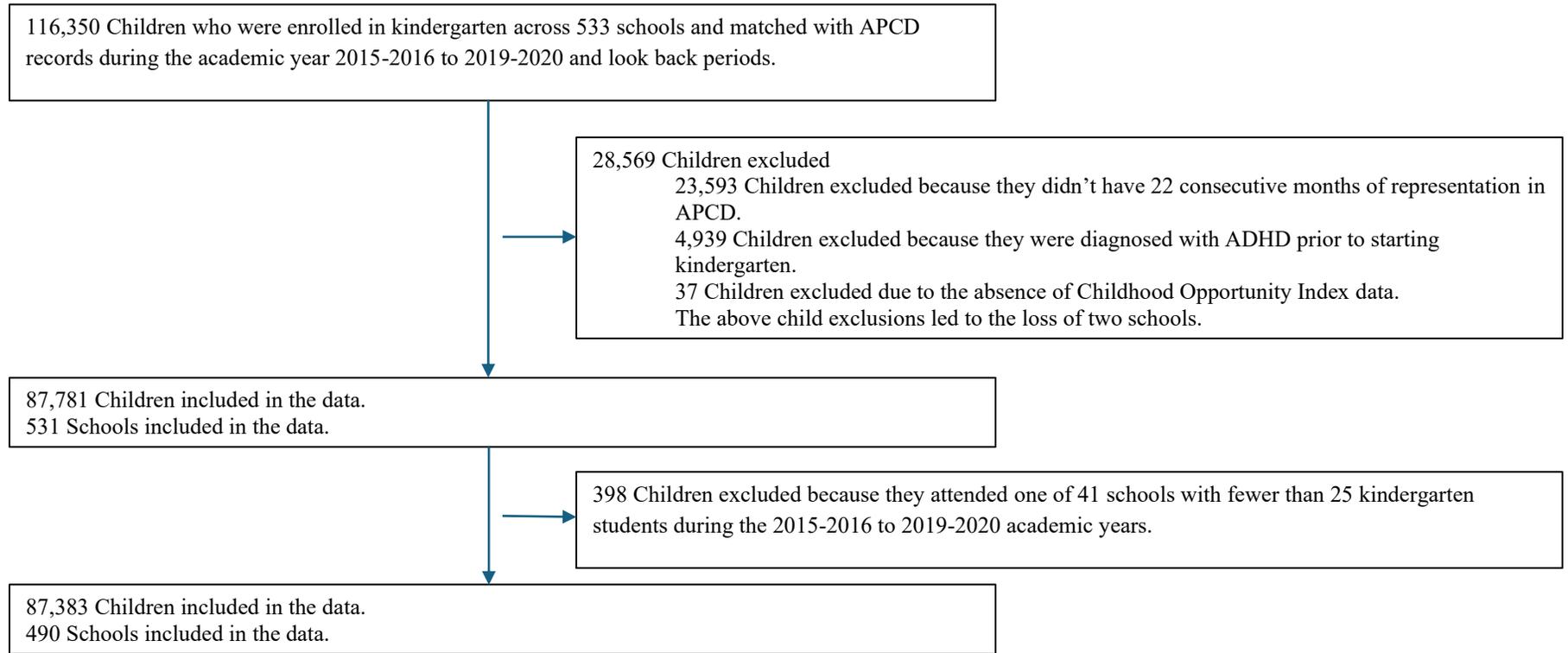
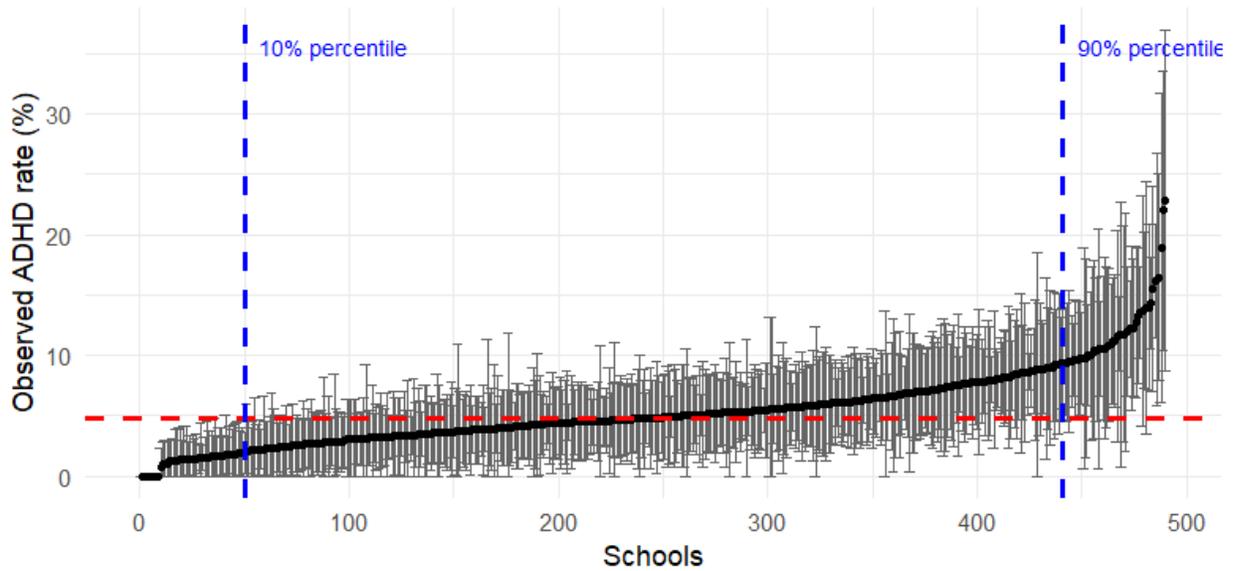
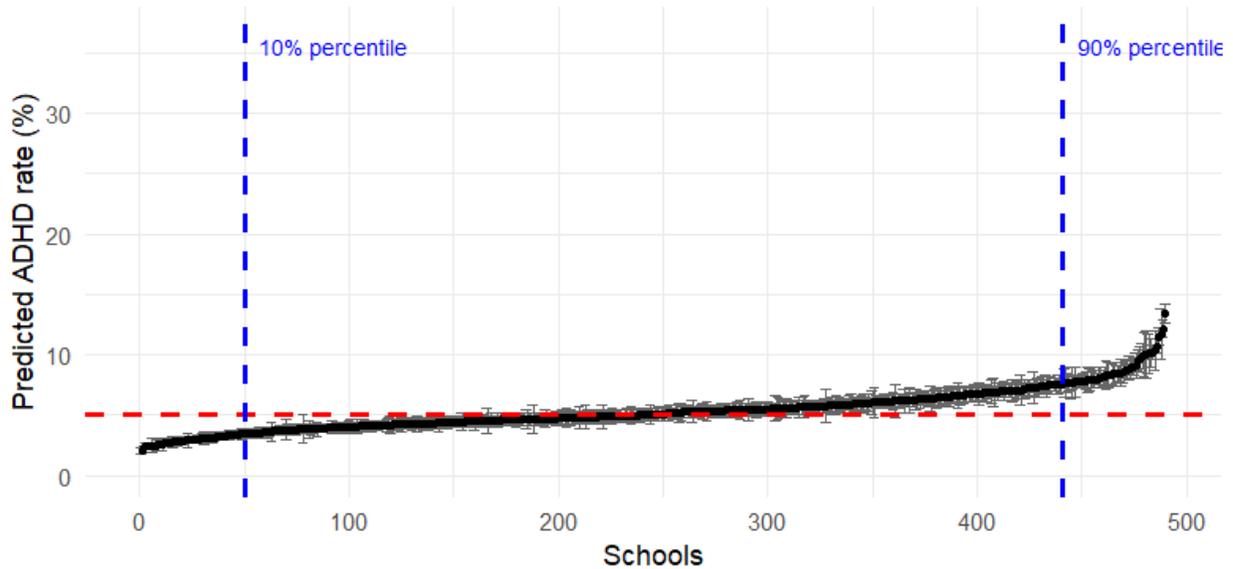


Figure 2. Unadjusted and model-adjusted rates of ADHD diagnosis.

Panel A.



Panel B.



Notes: Unadjusted and model-adjusted rates of ADHD diagnosis for 490 schools during study period with associated 95% confidence intervals, ordered from lowest to highest rate of ADHD diagnosis. (A) Unadjusted rates and (B) Model-adjusted rates. The median rate of ADHD is denoted in red. The median observed ADHD diagnosis rate shows a 4.9 times difference between the 10th and 90th percentiles. The median model-adjusted ADHD diagnosis rate shows a 2.2 times difference between the 10th and 90th percentiles.

Table 1. Sociodemographic of study population.

Characteristic	N / Mean N = 87,383	% / SD
ADHD diagnoses during kindergarten	4,520	5.2
Sex, n (%)		
Female	42,773	49.0
Male	44,610	51.0
Race/ethnicity, n (%).		
Hispanic	9,779	11.2
Non-Hispanic African American	20,021	22.9
Non-Hispanic White	52,240	59.8
Non-Hispanic other race	5,343	6.1
Other individual characteristics, n (%).		
Went to a school-based health center	2,130	2.4
Active in special education program	14,794	16.9
Changed schools during kindergarten	10,051	11.5
Young for grade	18,060	20.7
Individual level meal status, n (%).		
Meal status – free meal	59,699	68.3
Meal status – reduced price	6,707	7.7
Meal status - full-price paid	20,977	24.0
Co-occurring conditions of interest, n (%)		
Oppositional defiant disorder (ODD)	1,713	2.0
Other disruptive behavioral disorders	3,855	4.4
Autism (AUT)	1,635	1.9
Depression (DEP)	297	0.3
Anxiety (ANX)	3,229	3.7
Stress disorders	782	0.9
Specific learning disorder (SLD)	193	0.2
Tic disorder (TIC)	223	0.3
Used prescription stimulants during kindergarten	3,684	4.2
Received at least 1 disciplinary referral during kindergarten, n (%)	6,793	7.8
Region, n (%)		
Country	36,256	41.5
Delta	6,529	7.5
Mountain	4,743	5.4
Northwest	18,373	21.0
Suburban	10,521	12.0
Urban	10,961	12.5
Child opportunity scores, mean (sd)		
Overall	47.14	27.87
Education	47.98	27.85
Health and environment	47.81	28.02
Social and economic	47.14	27.97
School level meal delivery policies, n (%)		
Universal free meal	19,431	22.2
Breakfast after bell	29,851	34.2

Table 2. School characteristics by quintile of schools based on rates of children within the school newly diagnosed with ADHD during kindergarten

	Quintile of ADHD diagnoses rate during kindergarten						
	Q1	Q2	Q3	Q4	Q5	Q5/Q1	
Proportion of children newly diagnosed with ADHD (%)	1.7	3.4	4.6	6.1	10.0	5.71	+
Number of schools	98	98	98	98	98	NA	
Number of students	14171	16526	19399	20163	17124	NA	
Child characteristics aggregated to the school level (%)							
Female	49.2	49.3	48.6	49.0	48.3	0.98	
Male	50.9	50.7	51.4	51.0	51.7	1.02	
Hispanic	15.3	9.8	9.9	10.1	7.9	0.51	-
Non-Hispanic African American	12.5	19.6	22.4	28.4	35.5	2.83	+
Non-Hispanic White	64.5	64.7	62.2	56.3	52.0	0.81	-
Non-Hispanic other race	7.6	6.0	5.5	5.2	4.6	0.60	-
Went to a school-based health center	2.3	4.4	2.8	2.2	1.6	0.69	-
Active in special education program	14.9	17.4	17.7	17.1	17.9	1.20	+
Mobile	10.7	11.7	10.7	12.1	13.9	1.31	+
Young for grade	20.6	19.9	20.5	21.5	21.2	1.03	+
Meal status – free meal	65.6	65.5	65.7	73.0	77.0	1.17	+
Meal status – reduced price	8.3	8.2	8.1	7.0	6.5	0.79	-
Co-occurring conditions of interest (%)							
Oppositional defiant disorder (ODD)	1.2	1.3	1.7	2.1	3.2	2.76	+
Other disruptive behavioral disorders	2.4	3.4	4.1	5.2	6.7	2.73	+
Autism (AUT)	1.7	1.9	2.1	1.9	1.6	0.91	
Depression (DEP)	0.3	0.4	0.3	0.3	0.5	1.83	
Anxiety (ANX)	3.5	3.3	3.4	3.8	4.2	1.20	+
Post-traumatic stress disorder (PTSD)	0.8	0.9	0.9	0.9	1.1	1.34	+
Specific learning disorder (SLD)	0.3	0.2	0.2	0.2	0.2	0.73	
Tic disorder (TIC)	0.2	0.2	0.2	0.3	0.2	0.74	
Ever use of stimulants	1.8	3.0	3.8	4.9	7.9	4.52	+
Any disciplinary infraction (%)	5.6	6.5	7.0	7.9	11.0	1.94	+
Student Region (%)							
Country	29.9	35.8	33.4	34.3	48.4	1.62	+
Delta	2.0	5.9	6.1	11.9	13.2	6.47	+
Mountain	14.0	6.1	6.3	6.2	0.1	0.01	-
Northwest	38.7	24.5	20.5	16.9	5.2	0.13	-
Suburban	7.0	16.7	12.5	14.2	10.3	1.47	+
Urban	8.3	11.0	21.3	16.4	22.8	2.74	+
Child opportunity index (1-100)							
Overall COI	50.61	48.70	48.98	43.20	38.67	0.76	-
Education (ED)	49.27	46.25	48.06	46.03	39.76	0.81	-
Health and environment (HE)	52.80	47.66	49.35	43.86	38.42	0.73	-
Social and economic (SE)	50.97	49.78	49.06	42.94	39.34	0.77	-
School characteristics							
Total number of enrolled students	144.60	168.63	197.95	205.74	174.73	1.21	
Universal free meal (%)	16.9	19.2	22.0	30.0	31.5	1.86	+
Breakfast after bell (%)	31.8	31.3	37.9	38.0	30.8	0.97	

+/-: significant at 0.05 level based on the Cochran-Armitage test for trend (or the Jonckheere–Terpstra test for continuous variables).

Table 3. Adjusted odds ratios from mixed-effects logistic regression model for the outcome of ADHD diagnoses during kindergarten.

Factor	Odds ratio (95% CI)	P value
Sex (ref: Male)		
Female	0.39 (0.36 - 0.41)	<0.001
Race/ethnicity (ref: Non-Hispanic White)		
Hispanic	0.41 (0.36 - 0.47)	<0.001
Non-Hispanic African American	0.94 (0.86 - 1.03)	0.20
Non-Hispanic other race	0.67 (0.58 - 0.77)	<0.001
Other individual-level characteristics		
Went to a school-based health center	1.10 (0.89 - 1.35)	0.37
Mobile	1.50 (1.38 - 1.62)	<0.001
Young for grade	1.25 (1.16 - 1.34)	<0.001
Individual level meal status (ref: Full-price paid)		
Meal status – free meal	2.13 (1.94 - 2.34)	<0.001
Meal status – reduced price	1.67 (1.46 - 1.92)	<0.001
Child opportunity levels (ref: Very low)		
Education – low	1.02 (0.91 - 1.15)	0.68
Education – moderate	1.14 (1.00 - 1.29)	0.04
Education – high	1.14 (0.99 - 1.30)	0.07
Education – very high	1.15 (0.98 - 1.36)	0.10
Health and environment – low	0.94 (0.84 - 1.05)	0.26
Health and environment – moderate	0.90 (0.79 - 1.02)	0.11
Health and environment – high	0.90 (0.78 - 1.03)	0.13
Health and environment – very high	0.91 (0.78 - 1.06)	0.22
Social and economic – low	0.87 (0.78 - 0.97)	0.01
Social and economic – moderate	0.86 (0.76 - 0.97)	0.01
Social and economic – high	0.84 (0.74 - 0.96)	0.01
Social and economic – very high	0.78 (0.66 - 0.92)	0.003
School-level meal program		
Universal free meal	0.79 (0.71 - 0.88)	<0.001
Breakfast after bell	0.98 (0.91 - 1.07)	0.70
	Random effect (variance)	
School (Intercept)	0.11	
Student region (Intercept)	0.03	
Provider (Intercept)	0.05	
N	87,383	
N of schools	490	
N of providers	3,579	
N of regions	6	

Note: To facilitate model convergence, continuous COI levels were categorized as “very low”, “low”, “moderate”, “high”, and “very high”. Discrimination performance of this model as measured by area under the receiver operating characteristics curve was 0.722. Intraclass correlation coefficients (ICC) for the random effects terms were: school (ICC=0.032), medical provider (ICC = 0.015), and region (ICC=0.008).